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Quality, Credibility, and Utility:
The Relevance of Systematic Reviews



Symposium: Advances in Meta-analysis

Coordinator: Julio Sánchez Meca

Estimating an Overall Effect Size: Does the Between-Studies Variance Estimator have an Effect?

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<http://www.um.es/facpsi/metaanalysis>



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1. Three main objectives in meta-analysis

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- 1. To derive an average effect size through the integrated studies.**
 - 2. To test if the effect sizes are homogeneous around the average effect size.**
 - 3. If the effect sizes are heterogeneous, to look for the influence of moderator variables through statistical techniques as the analysis of variance or regression analysis.**

2. Procedure for averaging a set of independent effect sizes

- We focus on the standardized mean difference, d , as the effect-size index in a meta-analysis. Let d_i to be the effect size of the i study in a meta-analysis:

$$d = (m) \frac{\bar{Y}_{Ei} - \bar{Y}_{Ci}}{S}$$

- Assuming a random-effects model:

d_i estimates ~~μ~~

~~$$d_i = \mu + \tau_i + \epsilon_i$$~~

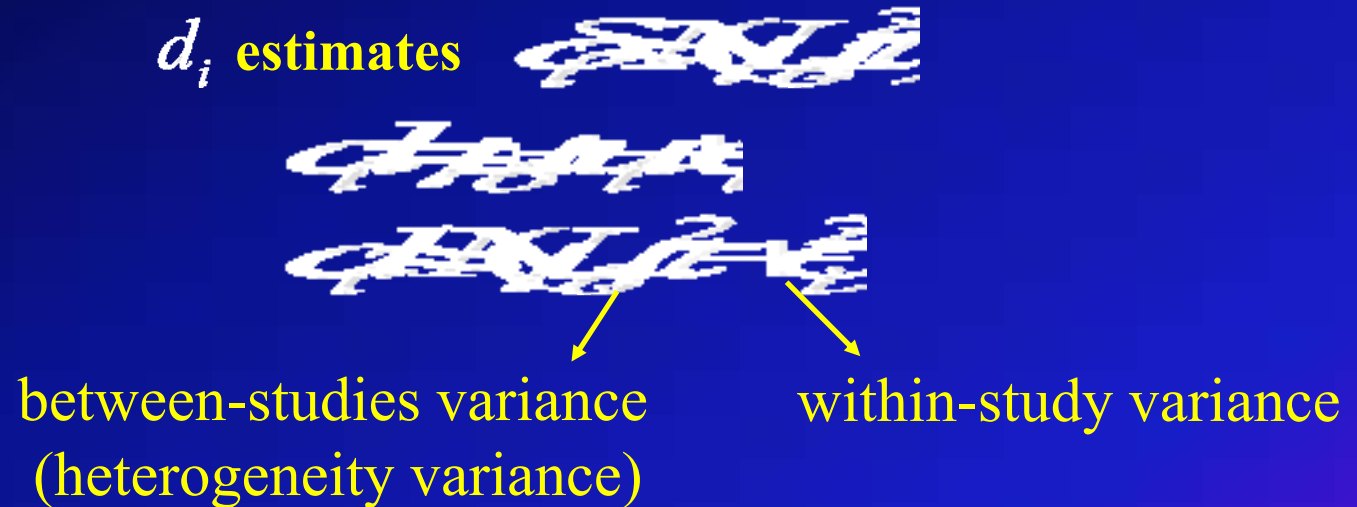
~~$$d_i = \mu + \tau_i + \epsilon_i$$~~

between-studies variance
(heterogeneity variance)

within-study variance

2. Procedure for averaging a set of independent effect sizes

➤ Assuming a random-effects model:



➤ General procedure for averaging:

$$\bar{d} = \frac{\sum w_i d_i}{\sum w_i}$$


heterogeneity variance estimator

The diagram shows a scatter plot of individual study effect sizes with error bars. An arrow points from the plot to the label 'heterogeneity variance estimator'.

3. Objective

➤ To compare the statistical performance (bias and mean squared error) of nine procedures to average a set of independent effect sizes based on different heterogeneity variance estimators.

$$\bar{d} = \frac{\sum w_i d_i}{\sum w_i}$$

$$w_i = \frac{1}{k_i}$$

- *Noniterative heterogeneity variance estimators:*

Hunter and Schmidt (1990):

$$\tau_{HS}^2 = \frac{Qk}{\sum w_i}$$

Hedges (1983):

$$\tau_{HE}^2 = \frac{\sum w_i (d_i - \bar{d})^2}{k-1} - \frac{1}{k} \sum \frac{d_i^2}{k_i}$$

3. Objective

$$\bar{d} = \frac{\sum w_i d_i}{\sum w_i}$$

~~$$\bar{d} = \frac{\sum w_i d_i}{\sum w_i}$$~~

- *Noniterative heterogeneity variance estimators:*

Dersimonian and Laird (1986): $\tau_{DL}^2 = \frac{Q - (k-1)}{c}$

Malzahn et al. (2000): $\tau_{MBH}^2 = \frac{\sum_i (1 - \phi_i) (d_i - \bar{d}_m)^2}{k-1} - \frac{1}{k} \sum_i \left(\frac{N_i}{n_m n_G} \right) - \frac{1}{k} \sum_i \phi_i d_i^2$

~~$$\tau_{MBH}^2 = \frac{\sum_i (1 - \phi_i) (d_i - \bar{d}_m)^2}{k-1} - \frac{1}{k} \sum_i \left(\frac{N_i}{n_m n_G} \right) - \frac{1}{k} \sum_i \phi_i d_i^2$$~~

3. Objective

$$\bar{d} = \frac{\sum w_i d_i}{\sum w_i}$$

~~$$\bar{d} = \frac{\sum w_i d_i}{\sum w_i}$$~~

- *Noniterative heterogeneity variance estimators:*

Hartum and Makambi (1986):

$$\hat{\tau}_{HM}^2 = \frac{\sigma^2}{[2(k-1) + \sigma^2]}$$

Sidik and Jonkman (2005):

$$\hat{\tau}_{SJ}^2 = \frac{\sum \hat{v}_i^{-1} (d_i - \bar{d}_\psi)^2}{k-1}$$

$$\hat{v}_i = r_i + 1$$

$$r_i = \sigma_i^2 / \sigma_0^2$$

$$\hat{\sigma}_0^2 = \frac{\sum (d_i - \bar{d}_\psi)^2}{k}$$

$$\bar{d}_\psi = \frac{\sum \hat{v}_i^{-1} d_i}{\sum \hat{v}_i^{-1}}$$

3. Objective

$$\bar{d} = \frac{\sum w_i d_i}{\sum w_i}$$

$$\bar{d} = \frac{\sum w_i d_i}{\sum w_i}$$

- *Iterative heterogeneity variance estimators:*

Maximum likelihood:

$$\tau_{ML}^2 = \frac{\sum w_i^2 [(d_i - \bar{d}_{ML}) - \delta_i^2]}{\sum w_i^2}$$

$$\bar{d}_{ML} = \frac{\sum w_i d_i}{\sum w_i} \quad w_i = 1/(\tau^2 + \delta_i^2)$$

Restricted maximum likelihood:

$$\tau_{RML}^2 = \frac{\sum w_i^2 [(d_i - \bar{d}_{ML}) - \delta_i^2]}{\sum w_i^2} + \frac{1}{\sum w_i}$$

3. Objective

$$\bar{d} = \frac{\sum w_i d_i}{\sum w_i}$$



- *Optimum average:*

$$\bar{d}_{\text{Optimum}} = \frac{\sum m d}{\sum m}$$



$$\bar{d}_{\text{Optimum}} = \frac{\sum w_i d_i}{\sum w_i}$$

4. Background

➤ Viechtbauer (2005) compared the statistical performance of five out of our nine procedures to average a set of independent effect sizes, using as the heterogeneity variance estimator: $\hat{\tau}_{HS}^2$ (Hunter & Schmidt, 1990), $\hat{\tau}_{HE}^2$ (Hedges 1983), $\hat{\tau}_{DL}^2$ (DerSimonian & Laird, 1986), $\hat{\tau}_{ML}^2$ (maximum likelihood estimator, Brockwell & Gordon, 2001), and $\hat{\tau}_{REML}^2$ (restricted maximum likelihood estimator, Brockwell & Gordon, 2001).

- **First conclusion:** Estimates of μ_δ were slightly negatively biased, the bias being the same regardless of which heterogeneity variance estimator was used to compute the average (p. 285).
- **Second conclusion:** Calculating the w_i values using unbiased estimates of τ^2 and σ_{di}^2 results in an estimate of the sampling variance of \bar{d} that is negatively biased (p. 263).



$$\bar{d} = \frac{\sum w_i d_i}{\sum w_i}$$

5. Method

- The effect-size index was the standardized mean difference, d_i :

$$d = (m) \frac{\bar{Y}_{Ei} - \bar{Y}_{Ci}}{s}$$

- With estimated within-study sampling variance, $\hat{\sigma}_i^2$:

$$d^2 = \frac{n_{Ei} + n_{Ci}}{n_{Ei} n_{Ci}} \frac{d^2}{2(n_{Ei} + n_{Ci})}$$

- A random-effects model was assumed, where $\delta_i \sim N(\mu_\delta, \tau^2)$:

$$\delta_i = \frac{\mu_{Ei} - \mu_{Ci}}{d}$$

- To define each parametric effect δ_i , two normal distributions, $N(\mu_{Ei}, \sigma_i^2)$ and $N(\mu_{Ci}, \sigma_i^2)$, were assumed

5. Method

- Pairs of independent samples with $n_E = n_C$ units were generated from these populations. Each pair of samples simulated the data of a primary study.
- The value of the parametric mean effect size was always set at $\mu_\delta = 0.5$
- Manipulated factors:
 - Variability of δ_i : $\tau^2 = 0, .04, .08, .16, \text{ and } .32$
 - Number of studies of the meta-analysis: $k = 5, 10, 20, 40, \text{ and } 100$
 - Average sample size of the studies in a meta-analysis: $\bar{N} = 30, 50, 80, 100$
- 10,000 replies (meta-analyses) for each of the 100 (5x4x5) conditions were simulated

5. Method

◆ Computations in each simulated meta-analysis:

- The standardized mean difference (d of Hedges) in each study
- The nine averages, \bar{d} , one for each τ^2 estimator:

\bar{d}_{HS} \bar{d}_{HE} \bar{d}_{DL} \bar{d}_{MBH} \bar{d}_{HM} \bar{d}_{SJ} \bar{d}_{ML} \bar{d}_{REML} \bar{d}_{OPTII}

$$\bar{d} = \frac{\sum w_i d_i}{\sum w_i}$$

~~$\bar{d} = \frac{\sum w_i d_i}{\sum w_i}$~~

5. Method

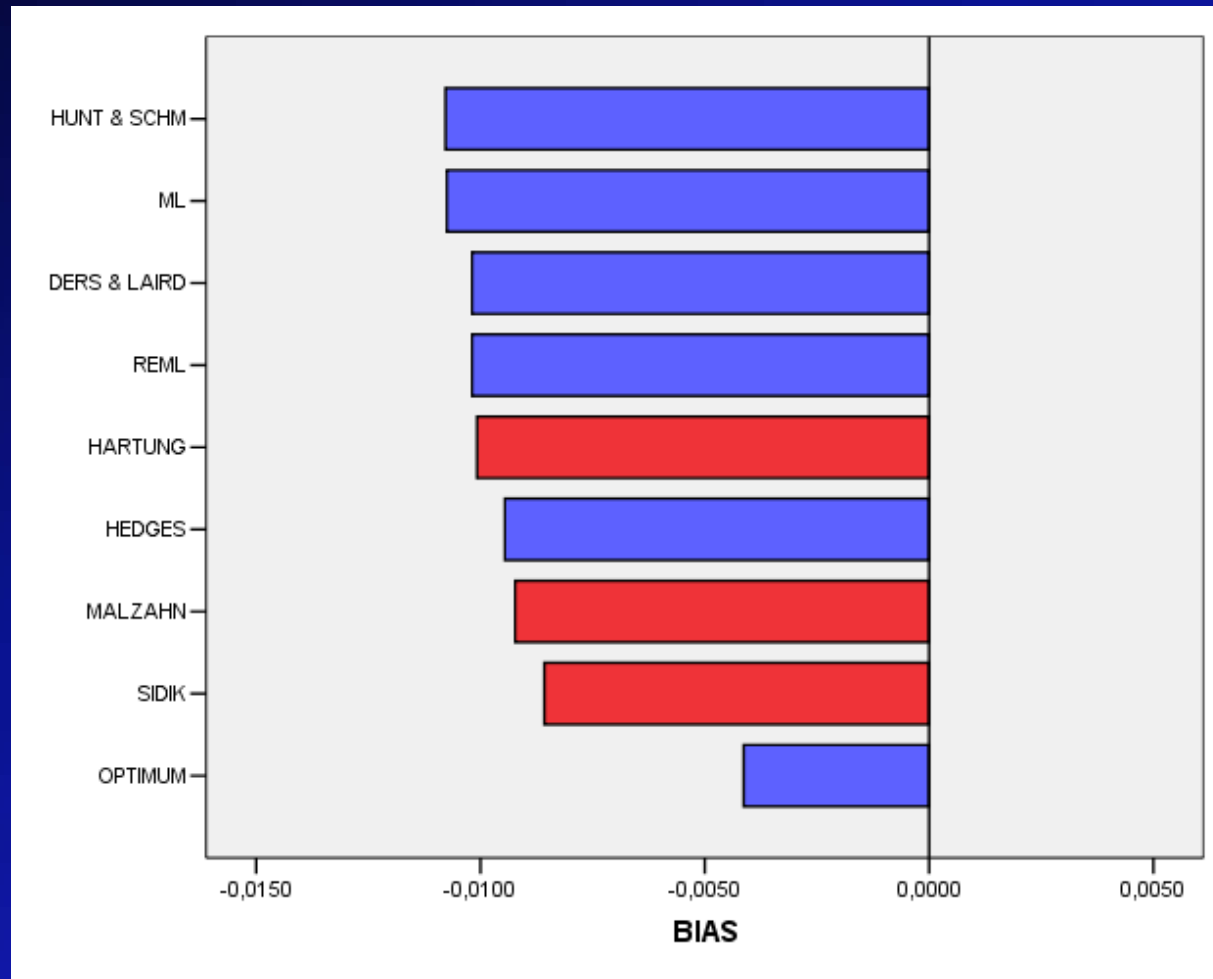
◆ Estimations through the 10,000 replies of a same meta-analysis

- Bias and Mean Squared Error of the nine averages:

$$\text{Bias}(\bar{d}) = \frac{\sum d_i}{10000} - \mu$$

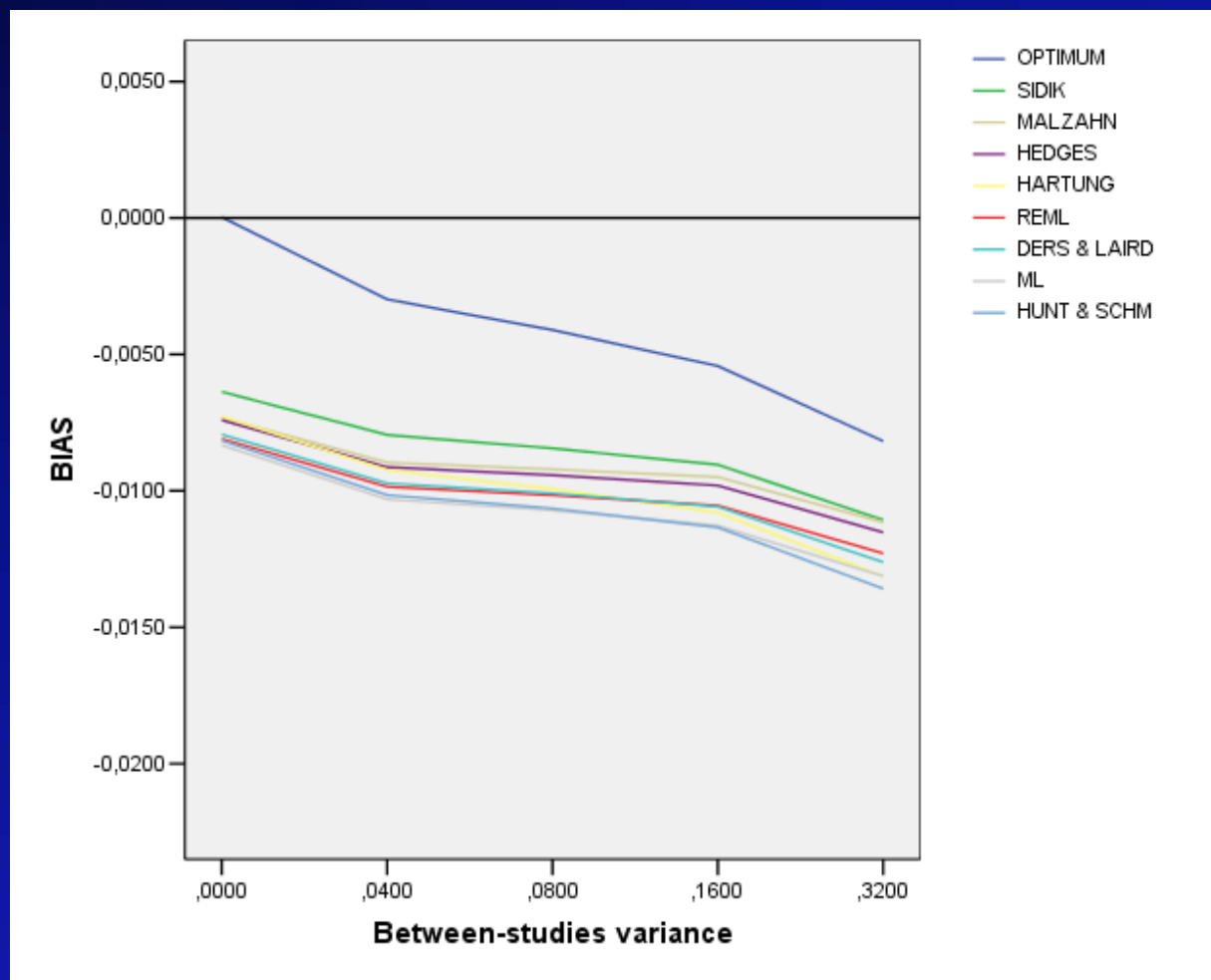
$$\text{MSE}(\bar{d}) = \frac{\sum (d_i - \mu)^2}{10000}$$

6. Results: BIAS of the nine averages

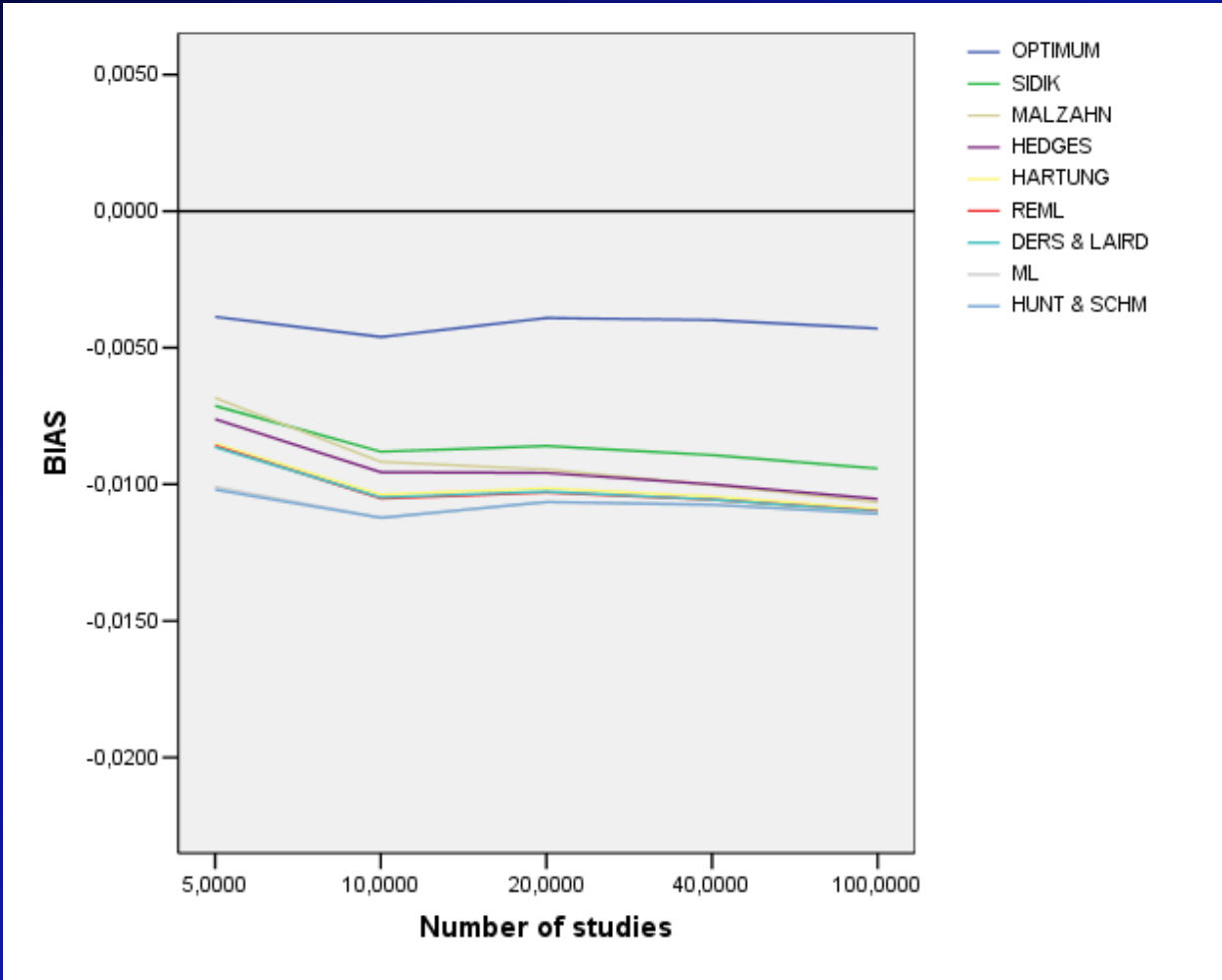


Sidik, Malzahn and Hedges procedures are the closest to the optimum one

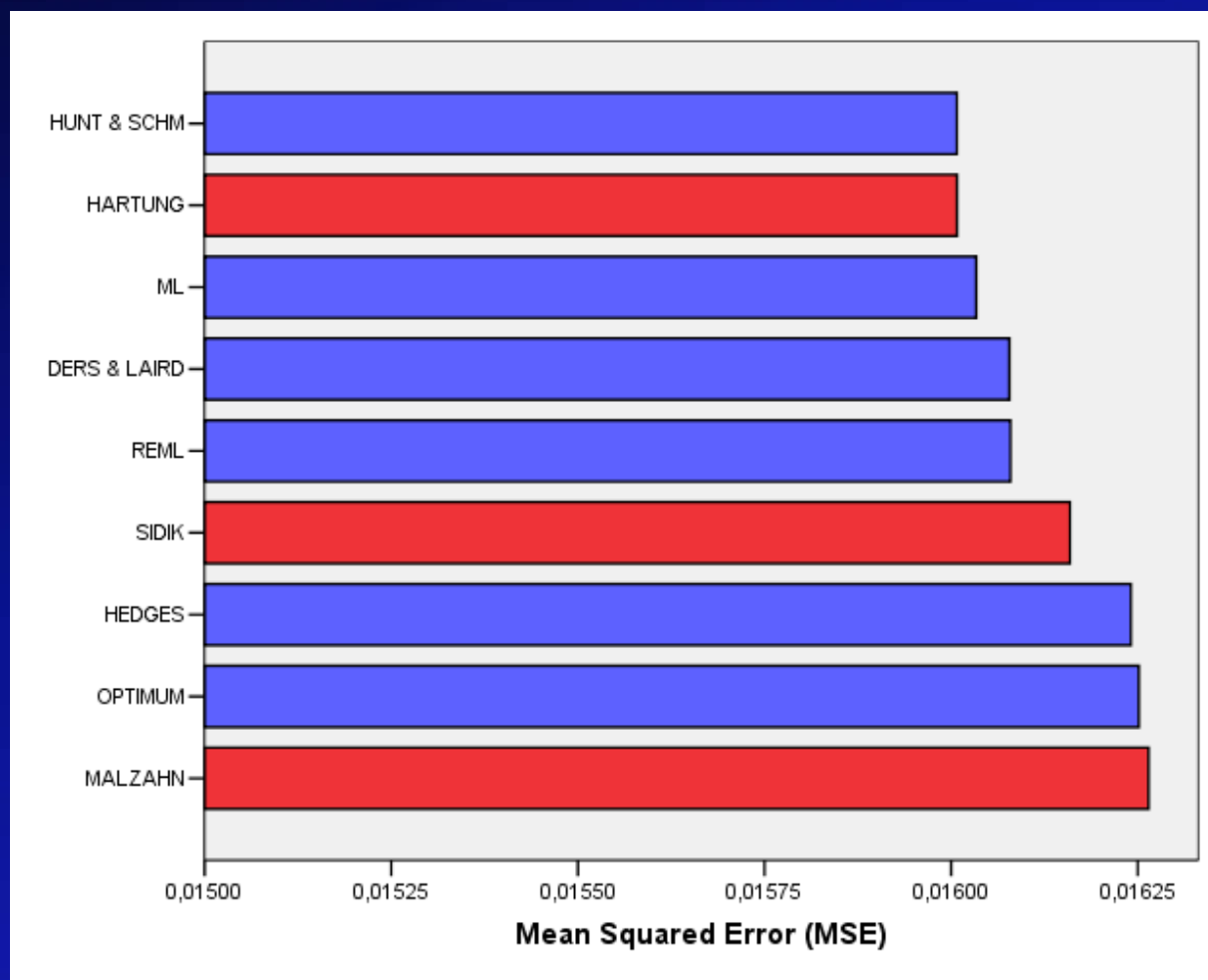
6. Results: BIAS as a function of the between-studies variance



6. Results: BIAS as a function of the number of studies

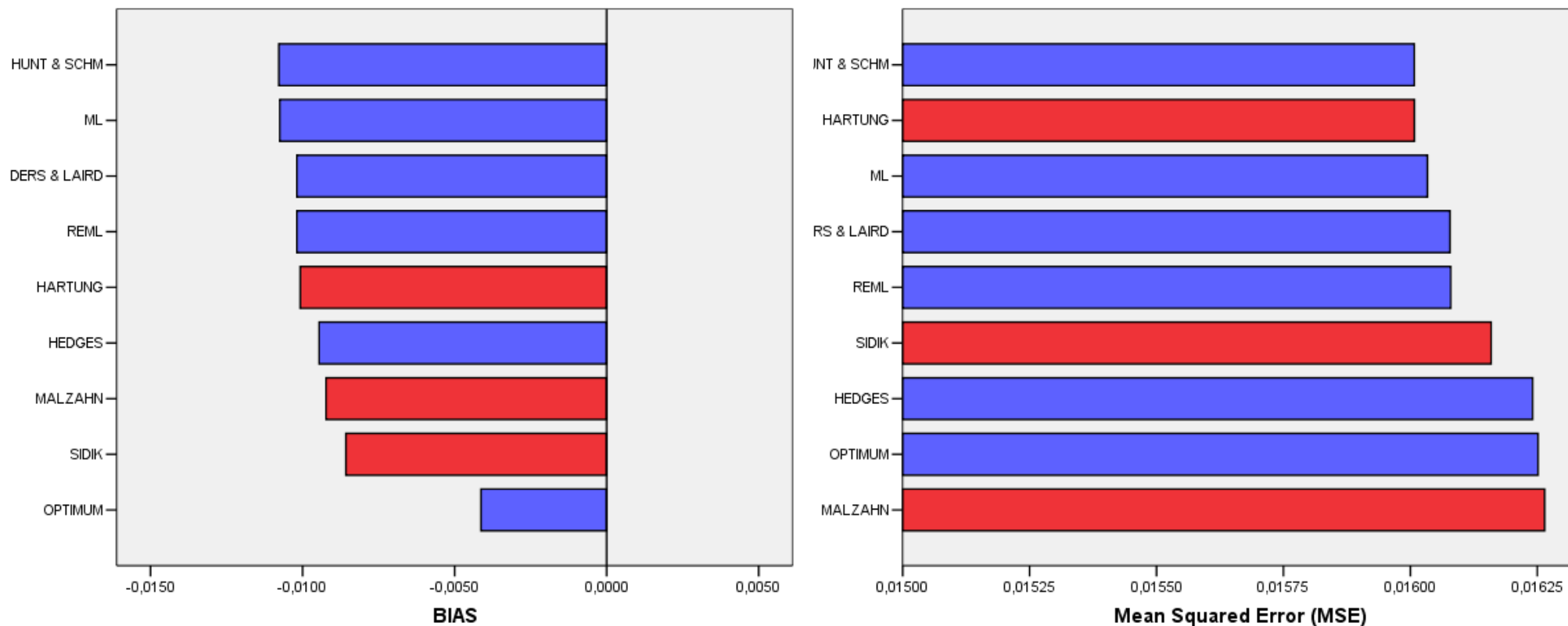


6. Results: Mean Squared Error of the nine averages



Sidik, Malzahn and Hedges procedures are the closest to the optimum one

6. Results: BIAS and MSE of the nine averages



Sidik, Malzahn and Hedges procedures are the closest to the optimum one

7. Conclusions

- 1. The selection among different heterogeneity variance estimators slightly affects to the bias and mean squared error of the overall effect size in a meta-analysis.**
- 2. All of the procedures to estimate an average effect size, with the different heterogeneity variance estimators, coincided in showing a slight negative bias. The largest bias values (around the 4%) were found in meta-analyses with small sample sizes and a large between-studies variance.**
- 3. With the exception of the Malzhan's procedure, the other ones slightly underestimated the mean squared error showed in the optimum average.**
- 4. Taking into consideration both the bias and the mean squared error, the procedures of Sidik & Jonkman (2005), Malzahn et al. (2000) and Hedges (1983) showed the best performance.**
- 5. The conclusions of our results are limited to the simulated conditions and the use of the standardized mean difference as the effect size index in a meta-analysis.**

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